# Week 7: Bayesian Inference and Parameter Estimation

Phong Le, Willem Zuidema

December 11, 2013

We are curious about some events or things (such as a language) and want to study their *hidden* mechanisms (grammar)  $G_{true}$ . A proper way to do is to collect a lot of data (sentences, dialogues)  $D = \{x_1, x_2, ..., x_n\}$  and then find a model  $\hat{G}$  that best *fits* (or explains) D. In this way, you expect that  $\hat{G}$  is a 'good' estimate of  $G_{true}$ .

In this lab, firstly, we will study one quality metric to measure the 'degree of belief' that a model G is a good estimate of  $G_{true}$  given observed data D: the posterior probability P(G|D), and how to compute it by using Bayesian inference. Then, we will examine two widely used estimation methods: Maximum Likelihood estimation (MLE) and Maximum A Posteriori estimation (MAP).

Required R Code At http://www.illc.uva.nl/LaCo/clas/fncm13/assignments/computerlab-week7/ you can find the R-files you need for this exercise.

### **1** Bayesian Inference

In statistics, according to Wikipedia, Bayesian inference

is a method of inference in which Bayes' rule is used to update the probability estimate for a hypothesis as additional evidence is acquired.

In other words, Bayesian inference is to compute the posterior probability P(G|D) based on the Bayes' rule

$$P(G|D) = \frac{P(D|G)P(G)}{P(D)}$$
(1)

where P(G) is the prior probability of G and D is additional evidence. In order to illustrate the method, let's examine the toy example below.

#### Toy Example: Murder in Dam Square

A man was found dead in Dam Square and two people, namely A and B, are suspected. After 24h investigating, the police found four witnesses, one of them reported that he saw A shooting the victim whereas the others said B. However, because it was foggy at that time, the police estimate that those witnesses only 80% correctly distinguished the two suspects. Our task is using Bayesian inference to help the police find out which one is the murderer, A or B.

First of all, we need to model the problem mathematically. Let's denote

- P(X) the prior probability that X is the murderer (note: P(X = B) = 1 P(X = A))
- $P(W_i|X)$  (i = 1..4) the confidence of the i-th witness' vision. Here,  $P(W_i = X|X) = 0.8$ .
- $P(X|W_{1,2,3,4})$  the posterior probability that X is the murderer based on the evidence given by all the four witnesses.

Our goal is to compute the posterior probability  $P(X = A | W_1 = A, W_2 = B, W_3 = B, W_4 = B)$  by updating the posterior probability when additional evidence is given as follows

- Step 0: when we don't have any evidence, we can only judge based on the prior probability P(X).
- Step 1: after the first witness reports, we update the posterior probability

$$P(X|W_1 = A) = \frac{P(W_1 = A|X)P(X)}{P(W_1 = A)}$$

where  $P(W_1 = A) = \sum_{X \in \{A,B\}} P(W_1 = A | X) P(X)$ .

Exercise 1.1: We set up the experiment as follows

p.prior = c(0.5, 0.5)1 # P(X=A) = P(X=B) = 0.5likelihood = matrix(c(0.8, 0.2, 0.2, 0.8), 2, 2) $\# P(W_i = X | X) = 0.8$ 2 3 # W1 = A, W2 = W3 = W4 = Bwitness = c(1, 2, 2, 2)

where we represent the likelihood-function as a matrix that gives for each actual killer (A,B) the likelihood of obtaining a witness-report incriminating A or B. Calculate (in R) the probability that A or B is the killer before and after hearing witness 1.

we then continue with incorporating the information from witnesses 2, 3 and 4. Note that the posterior after witness 1 becomes the prior for callulating the posterior after witness 2!

• Step 2: after the second witness reports, we update the posterior probability

$$P(X|W_1 = A, W_2 = B) = \frac{P(W_2 = B|X)P(X|W_1 = A)}{P(W_2 = B|W_1 = A)}$$

where  $P(X|W_1 = A)$  is computed in step 1. (Note: because  $W_i, W_j$  with  $i \neq j$  are independent given  $X, P(W_2 = B|X, W_1) = P(W_2 = B|X).)$ 

• Step 3: after the third witness reports, we update the posterior probability

$$P(X|W_1 = A, W_2 = B, W_3 = B) = \frac{P(W_3 = B|X)P(X|W_1 = A, W_2 = B)}{P(W_3 = B|W_1 = A, W_2 = B)}$$

where  $P(X|W_1 = A, W_2 = B)$  is computed in step 2.

• Step 4: after the last witness reports, we update the posterior probability

$$P(X|W_1 = A, W_2 = B, W_3 = B, W_4 = B) = \frac{P(W_4 = B|X)P(X|W_1 = A, W_2 = B, W_3 = B)}{P(W_4 = B|W_1 = A, W_2 = B, W_3 = B)}$$

where  $P(X|W_1 = A, W_2 = B, W_3 = B)$  is computed in step 3.

Exercise 1.2. The script murder. R automatizes the calculations at step 0-4.

• Step 0:

# step 01  $\mathbf{2}$ p.poste = p.prior3 print(p.poste)

• Step 1, 2, 3, 4: 1

 $\mathbf{2}$ 

```
 \begin{array}{c} \# \ step \ i > 0 \\ \textbf{for} \ (i \ in \ 1: \textbf{length}(witness)) \ \{ \\ \textbf{if} \ (witness [i] = 1) \ \# \ if \ the \ witness \ saw \ A \\ \end{array} 
3
 4
                              p.poste = p.poste * c(p.witness,1-p.witness)
```

else # if the witness saw B
 p.poste = p.poste \* c(1-p.witness,p.witness)
 p.poste = p.poste / sum(p.poste) # normalize
 print(p.poste)
}

Is the posterior probability at step 2 the same step 0? Explain why? Based on the posterior probability after step 4, who is the most suspected?

Exercise 1.3: In exercise 1, the prior distribution is uniform, because we haven't had any evidence yet. Now, assuming that B is a law-abiding citizen according to all records, whereas A has prior convictions for violence and other crimes. It might therefore be reasonable to suspect A more than B. We adjust the prior distribution as follows

1 p.prior = c(0.9, 0.1) # P(X=A) = 0.9, P(X=B) = 0.1

while keeping other parameters unchanged. Compute the posterior distribution as in exercise 1 and report what you get.

## 2 Parameter Estimation

 $\mathbf{5}$ 

6

7

8 9

10

In the previous section, we study how to use Bayesian inference to estimate a distribution. In this section, we will study how to select the 'best' model given observed data.

Maximum Likelihood Estimation (MLE) is a method to find values for model's parameters such that the likelihood given the observed data, e.g. the probability of the observed data given the model, is maximized

$$\hat{G}_{MLE} = \max_{G} P(D|G) \tag{2}$$

Maximum A Posteriori (MAP) Estimation on the other hand, is to maximize the posterior probability

$$\hat{G}_{MAP} = \max_{G} P(G|D) \tag{3}$$

According to the Bayes' theorem, we can compute posterior probability based on prior probability and likelihood, e.g.  $P(G|D) = \frac{P(D|G)P(G)}{P(D)}$ . Therefore

$$\hat{G}_{MAP} = \max_{G} \frac{P(D|G)P(G)}{P(D)} = \max_{G} P(D|G)P(G)$$

$$\tag{4}$$

(because P(D) is a constant in this case, we freely drop it).

In order to easily compute P(D|G) in Equation 2 and 4, observed data are assumed to be *independent* and *identically distributed* (i.i.d), e.g. examples are independently drawn from the same distribution. Hence

$$P(D = \{x_1, x_2, ..., x_n\}|G) = \prod_{i=1}^{n} P(x_i|G)$$
(5)

Exercise 2.1. What are the MLE and MAP hypotheses in exercise 1.3 after 4 witness reports? And what were they after the first 3 witness reports?

Because probabilities can become very small and multiplication is a relatively expensive operation, it is often convenient to work with the logarithm of probabilities.

Exercise 2.2. Confirm in R that :

$$\prod_{i} p_i = \exp\sum_{i} \log p_i$$

Now, Equation 2 and 4 respectively become  $^{1}$ 

$$\hat{G}_{MLE} = \max_{G} \prod_{i=1}^{n} P(x_i|G) = \max_{G} \sum_{i=1}^{n} \log P(x_i|G)$$
(6)

where the right hand side,  $\sum_{i=1}^{n} \log P(x_i|G)$ , is called *log-likelihood*, and

$$\hat{G}_{MAP} = \max_{G} P(G) \prod_{i=1}^{n} P(x_i|G) = \max_{G} \left( \log P(G) + \sum_{i=1}^{n} \log P(x_i|G) \right)$$
(7)

### Toy Example

In the following exercises, we will examine a very simple case: estimating the mean of a normal distribution  $N(x; \mu, \sigma^2)$ . The scenario is that, we draw a sample  $D = \{x_1, ..., x_n\}$  from  $N(x; \mu_{true}, \sigma_{true}^2)$ ; then, we ask you to estimate  $\mu_{true}$ . (Note that, in order to adapt the above equations, we need to replace probability by density.)

Note that, by the definition of a normal distribution, if x is distributed according to a normal distribution with mean  $\mu$  and standard deviation  $\sigma$  (i.e.,  $x \sim N(\mu, \sigma^2)$ ) then

$$p(x|\mu) = \frac{1}{2\sigma\sqrt{\pi}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$
(8)

which can be rewritten as

$$\log p(x|\mu) = -\frac{(x-\mu)^2}{2\sigma^2} + U$$
(9)

where U is a constant independent from  $\mu$  (and can often be, conveniently, ignored). Now, Equation 6 and 7 respectively become

$$\hat{\mu}_{MLE} = \max_{\mu} - \sum_{i=1}^{n} (x_i - \mu)^2 \tag{10}$$

$$\hat{\mu}_{MAP} = \max_{\mu} \left( \log p(\mu) - \sum_{i=1}^{n} (x_i - \mu)^2 \right)$$
(11)

Exercise 2.3: The file 'estimate\_mu.R' provides you with a visualization tool for the estimation problem (with both MLE and MAP): each time you press the Enter key, the program will draw an example from the true model and use it to update  $\hat{\mu}_{MLE}$  and  $\hat{\mu}_{MAP}$ ; after that, it will show a plot containing graphs of log-likelihood and log posterior probability over  $\mu$  and another plot containing graphs of  $\hat{\mu}_{MLE}$  and  $\hat{\mu}_{MAP}$  over sample size.

In this exercise, we assume that the prior distribution is also a normal distribution  $p(\mu) = N(\mu; \mu_{\mu}, \sigma_{\mu}^2)$ 

1. First of all, you need to set values for parameters and draw a sample

<sup>&</sup>lt;sup>1</sup>Note that because log is a monotonically increasing function,  $\max(a, b) = \max(\log(a), \log(b))$ .

```
1
    mu.true = 3
                        \# mean
    sigma.true = 10
                        \# standard deviation
2
3
    n = 100
                      # sample size
4
    data = rnorm(n, mean = mu.true, sd = sigmoid.true)
\mathbf{5}
6
                        # mean of mu (priori)
    mean.mu = 2.5
7
                      # standard deviation of mu (priori)
    sd.mu = 1
```

- 2. Before executing the file, try to predict how the graph of log-likelihood over  $\mu$  looks like, and how the graph of log-posterior-probability over  $\mu$  looks like when (i) observed data are ignored and (ii) observed data are used.
- 3. Load the file (source("estimate\_mu.R")), and then execute estimate.mu(data, sigma.true, mean.mu, sd.mu, plot=T) (note: the black lines are of MLE, the blue lines MAP). Report what you get.
- 4. It can be shown that  $\hat{\mu}_{MLE} = \frac{1}{n} \sum_{i=1}^{n} x_i$ . Confirm that by computing the sample average sum(data)/n. (Note:  $\hat{\mu}_{MLE}$  computed by the program is rounded.)
- 5. Change the prior p(μ) to have mean.mu = -2, sd.mu = 1 then execute estimate.mu(...) again. Now set mean.mu = -2, sd.mu = 1000 then execute estimate.mu(...). Do you have any conclusion about the effect of the prior distribution?

Exercise 2.4 (optional): In this exercise, we will compare MLE to MAP by computing mean squared errors over sample size.

1. First, we set up the experiment as in exercise 1

Then, we compute mean squared errors of m runs

```
mse.mle = \mathbf{rep}(0, n); mse.map = \mathbf{rep}(0, n)
1
\mathbf{2}
     m = 100
3
4
     for (i in 1:m) {
5
        data = rnorm(n, mean = mu.true, sd = sigma.true)
6
        mu.est = estimate.mu(data, sigma.true, mean.mu, sd.mu, plot=F)
7
         mse.mle = mse.mle + (mu.est mu.mle - mu.true)^2
8
         mse.map = mse.map + (mu.est mu.map - mu.true)^2
9
     }
10
     mse.mle = sqrt(mse.mle) / m
11
12
     mse.map = sqrt(mse.map) / m
```

And finally plot the errors

1 plot(1:n, mse.mle, type='l', ylim=c(min(min(mse.mle),min(mse.map)),max( max(mse.mle),max(mse.map))), xlab = 'sample size', ylab = 'MSE')
2 lines(1:n, mse.map, col='blue')

(Don't forget our notation: black is of MLE and blue MAP.)

2. Set n = 3000 and rerun the above.

3. Based on what you have done so far, draw conclusions about MLE vs MAP and when MAP is useful.

# 3 Submission

You have to submit a file named 'your\_name.pdf'. The deadline is 15:00 Monday 16 Dec. If you have any questions, contact Phong Le (p.le@uva.nl).